

Discussion of “Functional Models for Time-varying Random Objects”

Shahin TAVAKOLI
(Proposer)



October 23, 2019

Random Objects — Functional Data

Datapoints are

- 1 real, vector-valued, or Hilbert space valued curves
- 2 Covariance matrices/operators
- 3 measures/density functions
- 4 Trees

Abstract View

Fall into (roughly) 3 categories

- 1 Hilbert space (distance, directions, geometry)
- 2 Riemannian Manifold (distance, local Euclidean approximations \implies local directions, local geometry)
- 3 Metric space (only distance)

(local) geometry \implies we can do (local) moves in directions and (local) PCA

only distance \implies no notion of “moving into a direction”, **no PCA (?)**

Main contribution of Dubey & Müller

Data

Time-varying random objects, i.e. $X(t) \in \Omega, t \in [0, 1]$, where (Ω, d) is a *bounded* metric space.

Main Question

“We aim here at identifying dominant directions of variation [...] where random objects are indexed by time and in a general metric space”

How to define functional (object) PCA (fPCA); i.e.

- main modes of variations?
- (real-valued) scores for pair plots?
- estimation, consistency, examples

Key ideas

fPCA for $X(t) \in \mathbb{R}, t \in [0, 1]$ is based on **covariance kernel**
 $c(t, s) = \text{Cov}(X(t), X(s))$, so **covariance between random variables is key!**

Key observation (U, V random, U', V' iid copies)

If $U, V \in \mathbb{R}$ random, $d_E(U, V) = |U - V|$,

$$\begin{aligned}\mathbb{R} \ni \text{Cov}(U, V) &= \mathbb{E}[(U - \mathbb{E}U)(V - \mathbb{E}V)] \\ &= \frac{1}{4} \mathbb{E} [d_E^2(U, V') + d_E^2(U', V) - 2d_E^2(U, V)].\end{aligned}$$

If $U, V \in \mathbb{R}^p$ random, $d_E(U, V) = \sqrt{(U - V)^\top (U - V)}$,

$$\begin{aligned}\mathbb{R} \ni \text{Cov}_\Omega(U, V) &:= \frac{1}{4} \mathbb{E} [d_E^2(U, V') + d_E^2(U', V) - 2d_E^2(U, V)] \\ &= \mathbb{E} [(U - \mathbb{E}U)^\top (V - \mathbb{E}V)] \\ &\neq \text{Cov}(U, V) \in \mathbb{R}^{p \times p}\end{aligned}$$

Key ideas

If $U, V \in H$ random element of Hilbert space, $d_H(U, V) = \|U - V\|$,

$$\begin{aligned}\mathbb{R} \ni \text{Cov}^*(U, V) &:= \frac{1}{4} \mathbb{E} [d_H^2(U, V') + d_H^2(U', V) - 2d_H^2(U, V)]. \\ &= \mathbb{E} [\langle U - \mathbb{E}U, V - \mathbb{E}V \rangle] \\ &\neq \text{Cov}(U, V) \in S_1(H) \quad (\text{operators on } H)\end{aligned}$$

If $U, V \in H$ random object in metric space (Ω, d)

$$\begin{aligned}\mathbb{R} \ni \text{Cov}_\Omega(U, V) &:= \frac{1}{4} \mathbb{E} [d^2(U, V') + d^2(U', V) - 2d^2(U, V)]. \\ &\stackrel{?!}{=} \mathbb{E} [\langle U - \mathbb{E}U, V - \mathbb{E}V \rangle]\end{aligned}$$

Great! Now we can define $c(s, t) := \text{Cov}_\Omega(X(s), X(t))$... but positive definiteness is **not** guaranteed (hence no PCA...)

Metric spaces of Negative Types

Not all hope is lost: if (Ω, d) is of **negative type**, then there is a Hilbert space $(H, \langle \cdot, \cdot \rangle)$ and $f : \Omega \rightarrow H$ injective such that

$$d(U, V) = \langle f(U) - f(V), f(U) - f(V) \rangle.$$

(Sejdicinovic et al. 2013). Now we have

$$\begin{aligned} \mathbb{R} \ni \text{Cov}_\Omega(U, V) &:= \frac{1}{4} \mathbb{E} [d^2(U, V') + d^2(U', V) - 2d^2(U, V)]. \\ &= \mathbb{E} [\langle f(U) - \mathbb{E} f(U), f(V) - \mathbb{E} f(V) \rangle] \end{aligned}$$

The **blue bit** implies that

$$c(s, t) = \text{Cov}_\Omega(X(s), X(t)),$$

the **metric auto-covariance operator** is now symmetric positive semi-definite, so we are in business!

Functional (Object) PCA

Now usual fPCA realm!

Assume metric auto-covariance kernel $c(s, t) = \text{Cov}_\Omega(X(s), X(t))$ is **continuous** (typo in paper).

Mercer's Theorem:

$$c(s, t) = \sum_{k \geq 1} \lambda_k \phi_k(t) \phi_k(s),$$

uniform convergence (in t, s) and ϕ_k s are continuous.

Usually we then use the Karhunen–Loeve expansion

$$X(t) = \sum_{i \geq 1} \xi_i \phi_i(t); \quad \xi_i = \int X(t) \phi_i(t) dt$$

but not possible here!

Enters Fréchet

For usual functions $h(t) \in \mathbb{R}$, if $\int \phi(t)dt = 1$, we have

$$\int h(t)\phi(t)dt = \arg \inf_{\omega \in \mathbb{R}} \int d_E^2(\omega, h(t))\phi(t)dt.$$

Use this to define, for $S(t) \in \Omega$, the **generalized Fréchet integral**

$$\int_{\oplus} S(t)\phi(t)dt := \arg \inf_{\omega \in \Omega} \int d^2(\omega, S(t))\phi(t)dt$$

for $\int \phi = 1$.

→ need uniqueness and existence of infimum, hence **completeness** of the metric space (typo in paper).

Object FPC (“scores” in metric space)

Define **object FPC**

$$\psi_{\oplus}^k := \int_{\oplus} X \phi_k^* \in \Omega,$$

where $\phi_k^* = \phi_k / \int \phi_k$.

⇒ summarise $t \mapsto X(t) \in \Omega$ by a few $\psi_{\oplus}^k, k = 1, 2, 3$ (say).

⇒ We can plot these! (each one is in metric space)

How about scalar scores?

Letting $\mu_{\oplus}(t) := \arg \min_{\omega \in \Omega} \mathbb{E} [d^2(\omega, X(t))]$, define the k -th (scalar) **Fréchet score** by

$$\beta_k := \int d(X(t), \mu_{\oplus}(t)) \phi_k(t) dt \in \mathbb{R},$$

Now we can do pairs plots!

Remarks

We aim here at identifying dominant directions of variation. . . random objects are indexed by time and in a general metric space

So we wanted

- 1 General Metric Space
- 2 Dominant modes of variations (measuring co-variations and maximum variations)

Remarks (General Metric Space)

All examples considered have a structure richer than metric space:

- 1 densities
- 2 networks
- 3 covariances

If **only a metric** is available

\implies $\text{Cov}_\Omega(s, t)$ and its empirical eigenfunctions can be estimated, **but Object PC or Fréchet scores will not be computable easily** because

$$\mu_{\oplus}(t) := \arg \min_{\omega \in \Omega} \mathbb{E} [d^2(\omega, X(t))]$$

and

$$\psi_{\oplus}^k := \arg \inf_{\omega \in \Omega} \int d^2(\omega, X(t)) \phi_k^*(t) dt$$

involve minimizations, but **no gradient descent available!**

Remarks (measuring co-variations)

A lot of information is lost with metric covariance

If $\Omega = H$ a separable Hilbert space,

$$\begin{aligned}\mathbb{R} \ni \text{Cov}^*(U, V) &:= \frac{1}{4} \mathbb{E} [d_H^2(U, V') + d_H^2(U', V) - 2d_H^2(U, V)]. \\ &= \mathbb{E} [\langle U - \mathbb{E} U, V - \mathbb{E} V \rangle] \\ &= \text{Tr}(\text{Cov}(U, V)) \\ &\neq \text{Cov}(U, V) \in S_1(H) \quad (\text{operators on } H)\end{aligned}$$

Trace as measure of magnitude is

good for **self-adjoint positive semi-definite operator**,
but **not for general operators**.

Example (weakness of metric covariance)

Let $U(t) \in \mathbb{R}, t \in [0, 1/2]$ with mean zero and $\text{Cov}(U(t), U(t)) = 1$. Define $X(t) \in \mathbb{R}^2, t \in [0, 1]$ by

$$X(t) = \begin{cases} \begin{pmatrix} U(t) \\ 0 \end{pmatrix}, & t \in [0, 1/2], \\ \begin{pmatrix} 0 \\ U(t - 1/2) \end{pmatrix}, & t \in (1/2, 1]. \end{cases}$$

We have $\text{Cov}_\Omega(X(t), X(t + 1/2)) = 0$, but the true covariance matrix is

$$\text{Cov}(X(t), X(t + 1/2)) = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$$

and is non-zero. Hence metric covariance fails to measure association here. . .

Remarks (Modes of Variation)

fPCA is based on decomposition of covariance kernel/operator, but strong motivation comes from **variance maximization**:

In Hilbert space, find φ such that $\text{Var}(\langle \varphi, X \rangle)$ is maximized. If $X = X(t) \in H$ for each $t \in [0, 1]$, $\varphi(t) \in H$ and

$$\lambda_1 = \max_{\|\varphi\|=1} \text{Var}(\langle \varphi, X \rangle) = \int \left\langle \underbrace{\text{Cov}(X(s), X(t))}_{\text{operator on } H} \varphi(t), \varphi(s) \right\rangle ds dt$$

In the paper,

$$\lambda_1 = \max_{\int \phi^2 = 1} \int c(s, t) \phi(s) \phi(t) ds dt = \int \phi(s) \phi(t) \text{Tr}[\text{Cov}(X(s), X(t))] ds dt$$

fPCA Interpretation is lost if the space is actually a Hilbert space. . .

Fréchet scores

In the real-valued case ($\Omega = \mathbb{R}$),

$$\begin{aligned}\beta_k &:= \int d(X(t), \mu_{\oplus}(t)) \phi_k(t) dt \in \mathbb{R} \\ &= \int |X(t) - \mu(t)| \phi_k(t) dt\end{aligned}$$

This is **not the usual PC scores**, so interpretation is not straightforward. . .

Possible extensions

- 1 Instead of looking at $X(t), t \in [0, 1]$, we could have $t \in E$. For instance $E = \text{Great-Britain}$ and $X(t)$ is a covariance matrix at location $t \in E$. (Tavakoli et al. 2019)
- 2 Somehow use the implicit injection into Hilbert space to work with full covariance:

$$\begin{aligned}\mathbb{R} \ni \text{Cov}_\Omega(U, V) &:= \frac{1}{4} \mathbb{E} [d^2(U, V') + d^2(U', V) - 2d^2(U, V)]. \\ &= \mathbb{E} [\langle f(U) - \mathbb{E} f(U), f(V) - \mathbb{E} f(V) \rangle] \\ &= \text{Tr} \mathbb{E} [\underbrace{(f(U) - \mathbb{E} f(U)) \otimes (f(V) - \mathbb{E} f(V))}_{\text{full covariance}}]\end{aligned}$$

but not clear how to do since f is implicit...

Conclusions

Great thought provoking paper:

- moving beyond usual covariance (detects only linear dependencies),
 - hence away from Gaussian
- (Jim Ramsay's comments in SAMSI 2010, Diablerets 2016)

I propose the vote of thanks!